

# Statistics Ph.D. Qualifying Exam: Part I

October 7, 2022

Student Name: \_\_\_\_\_

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1. Answer 8 out of 12 problems. Mark the problems you selected in the following table.

|          |   |   |   |   |   |   |   |   |   |    |    |    |
|----------|---|---|---|---|---|---|---|---|---|----|----|----|
| Problem  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 |
| Selected |   |   |   |   |   |   |   |   |   |    |    |    |
| Scores   |   |   |   |   |   |   |   |   |   |    |    |    |

2. Write your answer right after each problem selected, attach more pages if necessary. Assemble your work in right order.
3. You can use the  $N(0,1)$  distribution table as attached.

1. Let  $Y$  be the sum of the observations of a random sample from a Poisson distribution with mean  $\theta$ . Let the prior pdf of  $\theta$  be gamma with parameters  $\alpha$  and  $\beta$ ,

$$f(\theta) = \frac{1}{\Gamma(\alpha)\beta^\alpha} \theta^{\alpha-1} e^{-\theta/\beta} \quad \alpha, \beta > 0.$$

- (a) Find the posterior pdf of  $\theta$ , given that  $Y = y$ .
- (b) If the loss function is  $[w(y) - \theta]^2$ , find the point estimate  $w(y)$  which minimize the loss function.

2. Let  $Y_1, Y_2, \dots, Y_n$  be i.i.d random variable from a distribution with probability density function,

$$f(y) = \frac{\theta^2}{1 + \theta} (1 + y) e^{-\theta y}$$

where  $y > 0$  and  $\theta > 0$ .

- (a) Find the maximum likelihood estimator (MLE) of  $\theta$ .
- (b) Find the MLE of  $1/\theta$ .

3. Let  $X_1, X_2, \dots, X_n$  be i.i.d random variable from a distribution with probability density function,

$$f(x|\theta) = \frac{2x}{\theta} e^{-x^2/\theta}$$

where  $x > 0$ .

- (a) Show that  $X_1^2$  is an unbiased estimator of  $\theta$ .
- (b) Find the Cramer-Rao lower bound for the variance of an unbiased estimator of  $\theta$ .
- (c) Find the uniformly minimum variance unbiased estimator (UMVUE) of  $\theta$ .

4. Let  $X$  and  $Y$  be independent random variables with  $X \sim \text{Exponential}(\lambda)$  and  $Y \sim \text{Exponential}(\mu)$ . Let  $Z$  and  $W$  be two random variables where  $Z = \min(X, Y)$  and

$$W = \begin{cases} 1 & \text{if } Z = X \\ 0 & \text{if } Z = Y \end{cases}$$

- (a) Find the joint distribution of  $Z$  and  $W$ .
- (b) Prove that  $Z$  and  $W$  are independent. (Hint: Show that  $P(Z \leq z | W = i) = P(Z \leq z)$  for  $i = 0$  or  $1$ .)

5. Consider the following model:  $(X_i, p_i)$  are independent pairs with

$$\begin{aligned} X_i | p_i &\sim \text{Bernoulli}(p_i), \\ p_i &\sim \text{Beta}(\alpha, \beta), \quad i = 1, \dots, n \end{aligned}$$

Let  $Y = X_1 + X_2 + \dots + X_n$ .

- (a) Find  $E(Y)$  and  $\text{Var}(Y)$ .
- (b) Find the marginal pdfs of  $X_i$ 's,  $i = 1, \dots, n$ , and use them to deduce the pdf of  $Y$ .

6. Let  $X_1, \dots, X_n$  be a random sample from the inverse Gaussian pdf

$$f(x|\mu, \lambda) = \sqrt{\frac{\lambda}{2\pi x^3}} \exp\left(-\frac{(\lambda - \mu)^2}{2\mu^2 x}\right), \quad x > 0$$

- (a) Find the method of moments estimator of  $\mu$  and  $\lambda$ . (Hint:  $E(X) = \mu$  and  $\text{Var}(X) = \frac{\mu^3}{\lambda}$ )
- (b) Find the MLE of  $\mu$  and  $\lambda$ .
- (c) Assume that  $\lambda$  is known. Is  $\hat{\mu}$ , the MLE for  $\mu$ , unbiased? Does the variance of  $\hat{\mu}$  attain the Rao-Cramer lower bound? Justify your answers.

7. Let  $U_1, U_2, U_3$  be iid Bernoulli with parameter  $\pi$ , that is

$$P(U_i = u_i) = \pi^{u_i}(1 - \pi)^{1-u_i} \quad u_i = 0, 1 \quad i = 1, 2, 3 \quad 0 < \pi < 1$$

Let  $W$  be independent of  $U_i$ ,  $i = 1, 2, 3$  with a Bernoulli distribution with parameter  $\theta$ , that is

$$P(W = w) = \theta^w(1 - \theta)^{1-w} \quad w = 0, 1 \quad 0 < \theta < 1.$$

Let  $X = WU_1 + (1 - W)U_2$  and  $Y = WU_1 + (1 - W)U_3$ .

- (a) Find the correlation between  $X$  and  $Y$ .
- (b) Find  $E(Y|X = x)$ .



8.  $X_1, X_2, \dots, X_n$  are iid with pdf given by

$$P(X = x; \theta) = (1 - \theta)\theta^{x-1} \quad x = 1, 2, \dots \quad 0 < \theta < 1.$$

Find the uniformly most powerful (UMP) test of size  $\alpha$  for testing  $H_0 : \theta = 1/3$  versus the alternative hypothesis  $H_1 : \theta > 1/3$ .

9.  $X_1, X_2, \dots, X_n$  is a random sample from density

$$f_X(s; \theta) = \theta e^{-\theta x} \quad x > 0 \quad \theta > 0.$$

- (a) Find the UMVUE of  $\theta$ .
- (b) Find the Cramer Rao lower bound of any unbiased estimator of  $\theta$ .
- (c) Find the mean squared error of the UMVUE.

10. Let  $X_1$  and  $X_2$  be two independent random variables, each with p.d.f  $f(x) = 1/2, 0 < x < 2$ . Therefore, the joint p.d.f. of  $X_1$  and  $X_2$  is

$$f(x_1, x_2) = 1/4, \quad 0 < x_1 < 2, 0 < x_2 < 2.$$

Let us consider

$$Y_1 = \frac{X_1}{X_1 + X_2}, \quad Y_2 = X_1 + X_2.$$

- (a) Find the joint pdf of  $Y_1$  and  $Y_2$ .  $g(x, y)$ .
- (b) Find the marginal p.d.f of  $Y_1$ ,  $g_1(y)$ .
- (c) Find the marginal p.d.f of  $Y_2$ ,  $g_2(y)$ . .

11. Let  $Y_1 < Y_2 < \cdots < Y_{10}$  be the order statistics of ten independent observations, say  $X_1, X_2, \cdots, X_{10}$ , from a standard uniform distribution,  $U(0, 1)$ . Let  $W$  be the number of these ten observations less than  $c$  ( $0 < c < 1$  is a constant).
- (a) What is the distribution of  $W$  ?
  - (b) What is the p.d.f. of  $Y_3$  ? What is its distribution ?
  - (c) Determine  $P(Y_3 < 0.6)$ .
  - (d) Find  $P(Y_5 < 0.6 < Y_8)$ .
  - (e) Find  $P(0.4 < Y_6 < 0.6)$ .

12. Let  $X_1, X_2$  be two independent random variables taking values from  $0, 1, 2, \dots, 9$  with equal probability.

- (a) Define the mgf of  $X_1$ .
- (b) Define the mgf of  $10X_2$ .
- (c) Let  $W = 10X_2 + X_1$ . Find the mgf of  $W$ .
- (d) Find the pmf of  $W$ .

Table of  $P(Z < z)$ ,  $Z \sim N(0,1)$ 

| <b>z</b>   | <b>0.00</b> | <b>0.01</b> | <b>0.02</b> | <b>0.03</b> | <b>0.04</b> | <b>0.05</b> | <b>0.06</b> | <b>0.07</b> | <b>0.08</b> | <b>0.09</b> |
|------------|-------------|-------------|-------------|-------------|-------------|-------------|-------------|-------------|-------------|-------------|
| <b>0.0</b> | 0.50000     | 0.50399     | 0.50798     | 0.51197     | 0.51595     | 0.51994     | 0.52392     | 0.52790     | 0.53188     | 0.53586     |
| <b>0.1</b> | 0.53983     | 0.54380     | 0.54776     | 0.55172     | 0.55567     | 0.55962     | 0.56356     | 0.56749     | 0.57142     | 0.57535     |
| <b>0.2</b> | 0.57926     | 0.58317     | 0.58706     | 0.59095     | 0.59483     | 0.59871     | 0.60257     | 0.60642     | 0.61026     | 0.61409     |
| <b>0.3</b> | 0.61791     | 0.62172     | 0.62552     | 0.62930     | 0.63307     | 0.63683     | 0.64058     | 0.64431     | 0.64803     | 0.65173     |
| <b>0.4</b> | 0.65542     | 0.65910     | 0.66276     | 0.66640     | 0.67003     | 0.67364     | 0.67724     | 0.68082     | 0.68439     | 0.68793     |
| <b>0.5</b> | 0.69146     | 0.69497     | 0.69847     | 0.70194     | 0.70540     | 0.70884     | 0.71226     | 0.71566     | 0.71904     | 0.72240     |
| <b>0.6</b> | 0.72575     | 0.72907     | 0.73237     | 0.73565     | 0.73891     | 0.74215     | 0.74537     | 0.74857     | 0.75175     | 0.75490     |
| <b>0.7</b> | 0.75804     | 0.76115     | 0.76424     | 0.76730     | 0.77035     | 0.77337     | 0.77637     | 0.77935     | 0.78230     | 0.78524     |
| <b>0.8</b> | 0.78814     | 0.79103     | 0.79389     | 0.79673     | 0.79955     | 0.80234     | 0.80511     | 0.80785     | 0.81057     | 0.81327     |
| <b>0.9</b> | 0.81594     | 0.81859     | 0.82121     | 0.82381     | 0.82639     | 0.82894     | 0.83147     | 0.83398     | 0.83646     | 0.83891     |
| <b>1.0</b> | 0.84134     | 0.84375     | 0.84614     | 0.84849     | 0.85083     | 0.85314     | 0.85543     | 0.85769     | 0.85993     | 0.86214     |
| <b>1.1</b> | 0.86433     | 0.86650     | 0.86864     | 0.87076     | 0.87286     | 0.87493     | 0.87698     | 0.87900     | 0.88100     | 0.88298     |
| <b>1.2</b> | 0.88493     | 0.88686     | 0.88877     | 0.89065     | 0.89251     | 0.89435     | 0.89617     | 0.89796     | 0.89973     | 0.90147     |
| <b>1.3</b> | 0.90320     | 0.90490     | 0.90658     | 0.90824     | 0.90988     | 0.91149     | 0.91309     | 0.91466     | 0.91621     | 0.91774     |
| <b>1.4</b> | 0.91924     | 0.92073     | 0.92220     | 0.92364     | 0.92507     | 0.92647     | 0.92785     | 0.92922     | 0.93056     | 0.93189     |
| <b>1.5</b> | 0.93319     | 0.93448     | 0.93574     | 0.93699     | 0.93822     | 0.93943     | 0.94062     | 0.94179     | 0.94295     | 0.94408     |
| <b>1.6</b> | 0.94520     | 0.94630     | 0.94738     | 0.94845     | 0.94950     | 0.95053     | 0.95154     | 0.95254     | 0.95352     | 0.95449     |
| <b>1.7</b> | 0.95543     | 0.95637     | 0.95728     | 0.95818     | 0.95907     | 0.95994     | 0.96080     | 0.96164     | 0.96246     | 0.96327     |
| <b>1.8</b> | 0.96407     | 0.96485     | 0.96562     | 0.96638     | 0.96712     | 0.96784     | 0.96856     | 0.96926     | 0.96995     | 0.97062     |
| <b>1.9</b> | 0.97128     | 0.97193     | 0.97257     | 0.97320     | 0.97381     | 0.97441     | 0.97500     | 0.97558     | 0.97615     | 0.97670     |
| <b>2.0</b> | 0.97725     | 0.97778     | 0.97831     | 0.97882     | 0.97932     | 0.97982     | 0.98030     | 0.98077     | 0.98124     | 0.98169     |
| <b>2.1</b> | 0.98214     | 0.98257     | 0.98300     | 0.98341     | 0.98382     | 0.98422     | 0.98461     | 0.98500     | 0.98537     | 0.98574     |
| <b>2.2</b> | 0.98610     | 0.98645     | 0.98679     | 0.98713     | 0.98745     | 0.98778     | 0.98809     | 0.98840     | 0.98870     | 0.98899     |
| <b>2.3</b> | 0.98928     | 0.98956     | 0.98983     | 0.99010     | 0.99036     | 0.99061     | 0.99086     | 0.99111     | 0.99134     | 0.99158     |
| <b>2.4</b> | 0.99180     | 0.99202     | 0.99224     | 0.99245     | 0.99266     | 0.99286     | 0.99305     | 0.99324     | 0.99343     | 0.99361     |
| <b>2.5</b> | 0.99379     | 0.99396     | 0.99413     | 0.99430     | 0.99446     | 0.99461     | 0.99477     | 0.99492     | 0.99506     | 0.99520     |
| <b>2.6</b> | 0.99534     | 0.99547     | 0.99560     | 0.99573     | 0.99585     | 0.99598     | 0.99609     | 0.99621     | 0.99632     | 0.99643     |
| <b>2.7</b> | 0.99653     | 0.99664     | 0.99674     | 0.99683     | 0.99693     | 0.99702     | 0.99711     | 0.99720     | 0.99728     | 0.99736     |
| <b>2.8</b> | 0.99744     | 0.99752     | 0.99760     | 0.99767     | 0.99774     | 0.99781     | 0.99788     | 0.99795     | 0.99801     | 0.99807     |
| <b>2.9</b> | 0.99813     | 0.99819     | 0.99825     | 0.99831     | 0.99836     | 0.99841     | 0.99846     | 0.99851     | 0.99856     | 0.99861     |
| <b>3.0</b> | 0.99865     | 0.99869     | 0.99874     | 0.99878     | 0.99882     | 0.99886     | 0.99889     | 0.99893     | 0.99896     | 0.99900     |
| <b>3.1</b> | 0.99903     | 0.99906     | 0.99910     | 0.99913     | 0.99916     | 0.99918     | 0.99921     | 0.99924     | 0.99926     | 0.99929     |
| <b>3.2</b> | 0.99931     | 0.99934     | 0.99936     | 0.99938     | 0.99940     | 0.99942     | 0.99944     | 0.99946     | 0.99948     | 0.99950     |
| <b>3.3</b> | 0.99952     | 0.99953     | 0.99955     | 0.99957     | 0.99958     | 0.99960     | 0.99961     | 0.99962     | 0.99964     | 0.99965     |
| <b>3.4</b> | 0.99966     | 0.99968     | 0.99969     | 0.99970     | 0.99971     | 0.99972     | 0.99973     | 0.99974     | 0.99975     | 0.99976     |
| <b>3.5</b> | 0.99977     | 0.99978     | 0.99978     | 0.99979     | 0.99980     | 0.99981     | 0.99981     | 0.99982     | 0.99983     | 0.99983     |
| <b>3.6</b> | 0.99984     | 0.99985     | 0.99985     | 0.99986     | 0.99986     | 0.99987     | 0.99987     | 0.99988     | 0.99988     | 0.99989     |
| <b>3.7</b> | 0.99989     | 0.99990     | 0.99990     | 0.99990     | 0.99991     | 0.99991     | 0.99992     | 0.99992     | 0.99992     | 0.99992     |
| <b>3.8</b> | 0.99993     | 0.99993     | 0.99993     | 0.99994     | 0.99994     | 0.99994     | 0.99994     | 0.99995     | 0.99995     | 0.99995     |
| <b>3.9</b> | 0.99995     | 0.99995     | 0.99996     | 0.99996     | 0.99996     | 0.99996     | 0.99996     | 0.99996     | 0.99997     | 0.99997     |
| <b>4.0</b> | 0.99997     | 0.99997     | 0.99997     | 0.99997     | 0.99997     | 0.99997     | 0.99998     | 0.99998     | 0.99998     | 0.99998     |
| <b>4.1</b> | 0.99998     | 0.99998     | 0.99998     | 0.99998     | 0.99998     | 0.99998     | 0.99998     | 0.99998     | 0.99999     | 0.99999     |